

APMA 1650 - Spring 2021

Lecture 11

Friday, Feb 12, 2021

Announcements

- ▶ Midterm Exam 1 Moved to Monday 2/22
- ▶ There will be a Homework 4 due 2/19
- ▶ TAs will hold review sessions next week (week 5)

Expectation

From last lecture we introduced the concept of the *expected value* of a discrete random variable.

Definition: Let X be a discrete random variable with range $R_X = \{x_1, x_2, x_3, \dots\}$. The **expected value** of X denoted by $EX = E(X) = E[X]$ is defined by

$$EX = \mu_X = \sum_{x_k \in R_X} x_k P_X(x_k)$$

$$P_X(x_k) \approx \frac{N_k}{N}$$

You should think of the expected value as the *average* or mean value that the random variable takes. μ_X is typically called the mean. **It is just a real (deterministic) number! (not random)**

Expectation:

Example: $X \sim \text{Bernoulli}(p)$. What is EX ?

$$R_x = \{0, 1\}$$

$$EX = 0 \cancel{(1-p)} + 1 p = p$$

0

Expectation:

Example: $X \sim \text{Geometric}(p)$. What is EX ?

$$\mathcal{R}_X = \{1, 2, \dots\} \quad P_X(k) = (1-p)^{k-1} p$$
$$k = 1, 2, 3, \dots$$

$$EX = \sum_{x_k \in \mathcal{R}_X} x_k P_X(x_k) = \sum_{k=1}^{\infty} k q^{k-1} p = p \sum_{k=1}^{\infty} k q^{k-1}$$

$$\sum_{k=0}^{\infty} q^k = \frac{1}{1-q}, \quad \frac{d}{dq} \sum_{k=0}^{\infty} q^k = \sum_{k=0}^{\infty} k q^{k-1} = \frac{1}{(1-q)^2}$$

$$\Rightarrow EX = p \frac{1}{(1-q)^2} = \frac{p}{p^2} = \boxed{1/p}$$

Expectation:

Example: $X \sim \text{Poisson}(\lambda)$. What is EX ?

$$\mathcal{R}_X = \{0, 1, 2, \dots\}, \quad P_X(k) = \frac{e^{-\lambda} \lambda^k}{k!}$$

$$EX = \sum_{x_k \in \mathcal{R}_X} x_k P_X(x_k) = \sum_{k=0}^{\infty} k \frac{e^{-\lambda} \lambda^k}{k!}$$

$$= e^{-\lambda} \sum_{k=1}^{\infty} \frac{\lambda^k}{(k-1)!} = \lambda e^{-\lambda} \sum_{k=1}^{\infty} \frac{\lambda^{k-1}}{(k-1)!}$$

$$= \lambda e^{-\lambda} \sum_{k=0}^{\infty} \frac{\lambda^k}{k!} = \lambda \frac{e^{-\lambda}}{e^{-\lambda}} = \boxed{\lambda}$$

Functions of random variables

Suppose X is a discrete random variable and $g : \mathbb{R} \rightarrow \mathbb{R}$, then $Y = g(X)$ is another random variable with range

$$R_Y = \{g(x) : x \in R_X\} = \{g(x_1), g(x_2), \dots\}.$$

What is the PMF of $Y = g(X)$?

$$P_Y(y) = P(g(X) = y) = \sum_{\{x : g(x)=y\}} P_X(x)$$

Sum over all the probabilities of values $x \in R_X$ that lead to $g(x) = y$, can have multiple values if g is not one-to-one.

Functions of random variables

Example X is a discrete random variable with $P_X(k) = \frac{1}{5}$ for $k \in \{-2, -1, 0, 1, 2\}$.

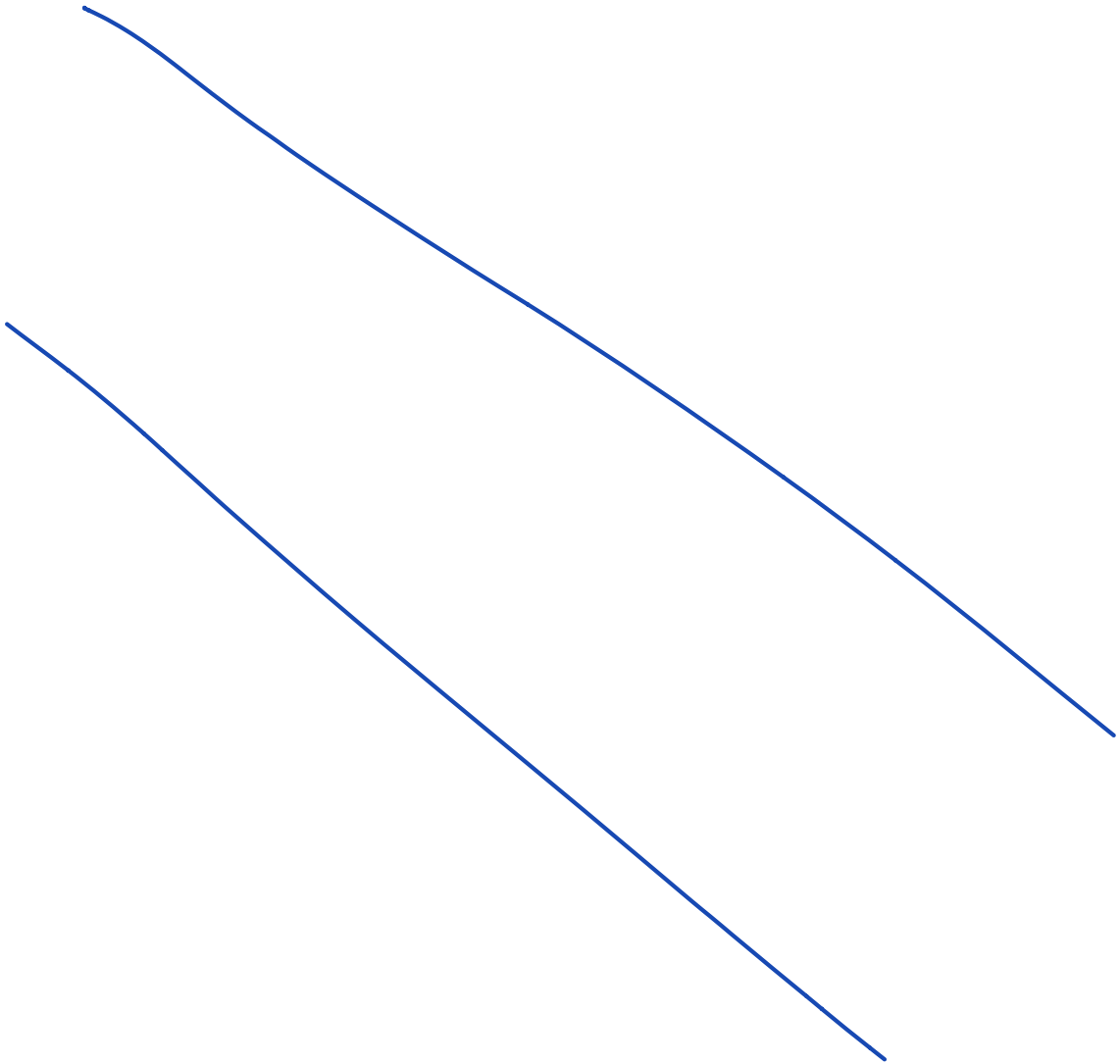
- ▶ What is $P_Y(y)$ for $Y = |X|$?
- ▶ What is $E|X|$?

$Y = X $	0	1	2
$P_Y(y)$	$\frac{1}{5}$	$\frac{2}{5}$	$\frac{2}{5}$

$$E|X| = 0 \cdot \frac{1}{5} + 1 \cdot \frac{2}{5} + 2 \cdot \frac{2}{5} = \frac{6}{5}$$

$\wedge \qquad \qquad \wedge$

$1 \cdot \frac{1}{5} + 1 \cdot 1 \cdot \frac{1}{5} \qquad 2 \cdot \frac{1}{5} + 1 \cdot 2 \cdot \frac{1}{5}$



Functions of Random Variables

We can calculate $E[g(X)]$ without having to know the PMF of $g(X)$

Law of the unconscious statistician (LOTUS) Let X be a discrete RV and let $g : \mathbb{R} \rightarrow \mathbb{R}$, then

$$E[g(X)] = \sum_{x_k \in R_X} g(x_k) P_X(x_k)$$

This law is "obvious" to a statistician. Think in terms of relative frequencies...

$$\text{average of } g(X) = \frac{1}{N} \sum_{x_k \in R_X} \{\# \text{ of times } x_k \text{ occurs}\} g(x_k)$$

Proof

$$\begin{aligned} E[Y] &= \sum_{y \in R_Y} y P_Y(y) \\ &= \sum_{y \in R_Y} y \sum_{x: g(x)=y} P_X(x) \\ &= \sum_{y \in R_Y} \sum_{x: g(x)=y} g(x) P_X(x) \\ &= \sum_{x \in R_X} g(x) P_X(x) \end{aligned}$$

Here we used the partition

$$R_X = \bigcup_{y \in R_Y} \{x : g(x) = y\}$$

Properties of Expectation

Linearity of Expectation: Let $a, b \in \mathbb{R}$ and X be a discrete RV

$$E[aX + b] = aE[X] + b$$

$\mu_Y = a\mu_X + b$

Proof: Take $g(x) = ax + b$, then

$$\begin{aligned} E[g(X)] &= \sum_{x \in R_X} (ax + b)P_X(x) \\ &= a \underbrace{\sum_{x \in R_X} xP_X(x)}_{EX} + b \underbrace{\sum_{x \in R_X} P_X(x)}_{=1} \end{aligned}$$

Properties of Expectation

More Generally: Let X_1, \dots, X_n be ANY discrete random variables

$$E[X_1 + X_2 + \dots + X_n] = EX_1 + EX_2 + \dots + EX_n$$

Note there is no assumption of independence here

We need the concept of **joint probability** to prove this (this will be later)

Expectation:

Example: $X \sim \text{Binomial}(n, p)$. What is EX ?

$$EX = \sum_{k=0}^n k \binom{n}{k} p^k (1-p)^{n-k} = ?$$

$$X = X_1 + X_2 + \dots + X_n \quad \leadsto \quad X_i \sim \text{Bernoulli}(p)$$

$$EX = \sum_{k=0}^n EX_i = \sum_{k=0}^n p = \underline{\underline{np}}$$

↙ Geometric(p).

How about $X \sim \text{Pascal}(m, p)$?

$$X = X_1 + X_2 + \dots + X_m$$

Expectation

Example: n students turn in homework to m dropboxes. Each student picks a dropbox at random. What is the expected number of empty dropboxes?

$$X_i = \begin{cases} 1 & \text{if box } i \text{ is empty} \\ 0 & \text{otherwise} \end{cases}$$

$$X = \sum_{i=1}^m X_i = \# \text{ of empty dropboxes.}$$

$$\mathbb{E} X_i = \text{prob box } i \text{ empty} = \left(\frac{m-1}{m}\right)^n$$
$$\mathbb{E} X = \sum_{i=1}^m \mathbb{E} X_i = m \left(\frac{m-1}{m}\right)^n$$

Variance

A measure of how spread out a random variable (or PMF) is about its mean $\mu_X = EX$

how far from mean

$$E[\overbrace{X - \mu_X}] = E[X] - \mu_X = \mu_X - \mu_X = 0.$$

Definition Let X be a random variable with mean $\mu_X = EX$, then the **variance** of X is defined by

$$\text{Var}(X) = E[(X - \mu_X)^2]$$

The **standard deviation** is defined by

$$\text{SD}_X = \sigma_X = \sqrt{\text{Var}(X)}$$

A useful formula

Computational formula: Let X be a random variable with mean $\mu_X = EX$,

$$\text{Var}(X) = E[X^2] - \mu_X^2$$

Proof:

$$\begin{aligned}\text{Var}(X) &= E(X - \mu_X)^2 \\ &= E[X^2 - 2\mu_X X + \mu_X^2] \\ &= E[X^2] - 2\mu_X EX + \mu_X^2 \\ &= E[X^2] - 2\mu_X^2 + \mu_X^2 \\ &= E[X^2] - \mu_X^2.\end{aligned}$$

Variance

Example: $X \sim \text{Bernoulli}(p)$. What is $\text{Var}(X)$?

$$\text{Var}(X) = EX^2 - \mu_X^2$$

$$= \cancel{0^2(1-p)} + 1^2 p - p^2$$

$$= p^2 - p = p(1-p) = pq.$$

Properties of Variance

Rescaling: Let X be a random variable and $a, b \in \mathbb{R}$, then

$$\text{Var}(aX + b) = a^2\text{Var}(X)$$

Proof: Let $Y = aX + b$. Note that $\mu_Y = a\mu_X + b$, so

$$Y - \mu_Y = (aX + b) - (a\mu_X + b) = a(X - \mu_X)$$

Therefore

$$\text{Var}(Y) = E[(Y - \mu_Y)^2] = E[a^2(X - \mu_X)^2] = a^2\text{Var}(X)$$

Shifting a random variable by a constant doesn't change its variance. Rescaling

Properties of Variance

More generally: Let X_1, X_2, \dots, X_n be **independent** random variables, then

$$\text{Var}(X_1 + X_2 + \dots + X_n) = \text{Var}(X_1) + \text{Var}(X_2) + \dots + \text{Var}(X_n).$$

We will be able to show this later using the idea of **covariance**.

This is not true if the random variables are not independent!

Variance

Example: $X \sim \text{Binomial}(n, p)$. What is $\text{Var}(X)$?

$$X = X_1 + X_2 + \dots + X_n, \quad X_i \sim \text{Bernoulli}(p) \\ \text{independent.}$$

$$\text{Var}(X) = \sum_{k=1}^n \text{Var}(X_k) = \underline{\underline{n p (1-p)}}.$$