

# Making Estimators Unbiased.

Recall: Given a random sample  $X_1, X_2, \dots, X_n$   
an estimator  $\hat{\theta}$  for  $\theta$ .

$$\text{Bias: } B(\hat{\theta}) = E\hat{\theta} - \theta$$

Can we make a biased estimator unbiased?

Suppose  $B(\hat{\theta}) \neq 0$  - biased.

Create new estimator  $\hat{\theta}_1 = c \hat{\theta}$

$$\begin{aligned} B(\hat{\theta}_1) &= E[c\hat{\theta}] - \theta \\ &= c E\hat{\theta} - \theta \\ &= c [B(\hat{\theta}) + \theta] - \theta = 0 \end{aligned}$$

what is  $c$ ?

$$c = \frac{\theta}{B(\hat{\theta}) + \theta}$$

$$\hat{\theta}_1 = \frac{\theta}{B(\hat{\theta}) + \theta} \hat{\theta} \rightarrow \text{unbiased.}$$

Ex Sample Variance.

$$X_1, X_2, X_3, \dots, X_n$$

$$\bar{S}_n = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2$$

$$= \frac{1}{n} \sum_{i=1}^n X_i^2 - (\bar{X})^2$$

$$\mathbb{E} \bar{S}_n = \mathbb{E} X_i^2 - \mathbb{E} (\bar{X})^2$$

$$= \text{Var}(X_i) + \cancel{\mu^2} - \text{Var}(\bar{X}) - \cancel{\mu^2}$$

$$= \text{Var}(X_i) - \text{Var}(\bar{X})$$

$$= \sigma^2 - \frac{\sigma^2}{n}$$

$$\Rightarrow B(\bar{S}_n) = -\frac{\sigma^2}{n}$$

↙ Biased.

Make unbiased.

$$C = \frac{0}{B(\bar{S}_n) + 0} = \frac{\sigma^2}{-\frac{\sigma^2}{n} + \sigma^2} = \frac{n}{n-1}$$

Unbiased estimator for the variance.

$$S^2 = \frac{n}{n-1} \bar{S}_n = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$$

# Maximum likelihood estimation.

Ex Container with 3 balls (red or white).

Goal: Estimate  $\theta = \#$  of red balls.

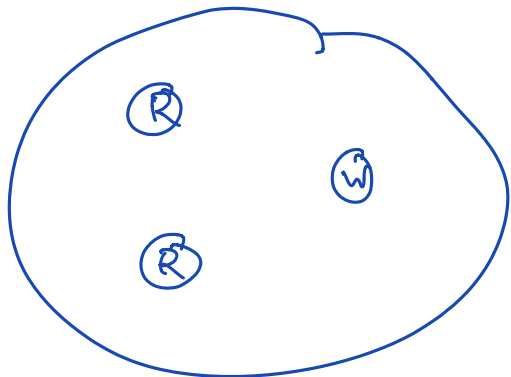
Sampling twice (w/o replacement).

↳ get two reds.

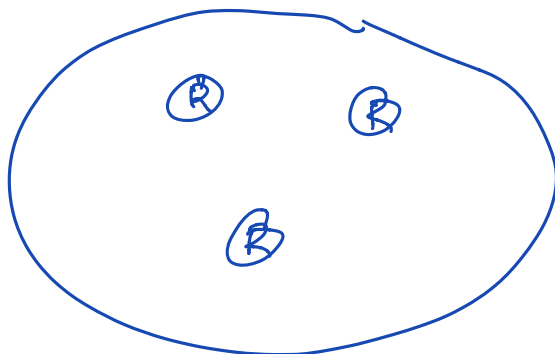
What is the best estimate for  $\theta$ ?

Two options ..

$$\theta = 2$$



$$\theta = 3$$



$$P_{\theta=2}(2 \text{ reds}) = \frac{1}{\binom{3}{2}} = \frac{1}{3} \quad \Bigg| \quad P_{\theta=3}(2 \text{ reds}) = ($$

$\Rightarrow \theta = \frac{1}{3}$  is the most likely.

---

Ex Sampling is w/ replacement.

↳ Samples are independent.

Sample  $n = 4$  times,  $X_1, X_2, X_3, X_4$

$X_i = \begin{cases} 1 & i^{\text{th}} \text{ sample red} \\ 0 & i^{\text{th}} \text{ sample white} \end{cases}$

$X_i \sim \text{Bernoulli}(\frac{\theta}{3})$

Suppose  $X_1 = 1, X_2 = 0, X_3 = 1, X_4 = 1$

$(1, 0, 1, 1)$

What is the most likely  $\theta$ ?

Likelihood  $P(x_1, x_2, x_3, x_4; \theta)$   
|  $\theta$  parameter.  
 $= P(x_1; \theta) P(x_2; \theta) P(x_3; \theta) P(x_4; \theta)$

↳ independence.

$$P(1, 0, 1, 1; \theta) = \frac{\theta}{3} \left(1 - \frac{\theta}{3}\right) \frac{\theta}{3} \frac{\theta}{3} = \left(\frac{\theta}{3}\right)^3 \left(1 - \frac{\theta}{3}\right)$$

$\theta$	0	1	2	3
$P(1, 0, 1, 1; \theta)$	0	0.0747	0.0988	0

↑ biggest.

$\theta = 2$  is the most likely.

General Method.

$$X_1, X_2, \dots, X_n$$

random sample from a distribution with param.  $\theta$ .

Def

$$L(x_1, x_2, \dots, x_n; \theta) = \begin{cases} P_{X_1, X_2, \dots, X_n}(x_1, x_2, \dots, x_n; \theta) & \downarrow \text{discrete} \\ f_{X_1, X_2, \dots, X_n}(x_1, x_2, \dots, x_n; \theta) & \uparrow \text{continuous} \end{cases}$$

Maximum Likelihood Estimator (MLE).

A maximum likelihood estimator

$$\hat{\theta}_{ML} = \hat{\theta}_{ML}(x_1, x_2, \dots, x_n)$$

is a value of  $\theta$  that maximizes

$$\theta \mapsto L(x_1, x_2, \dots, x_n; \theta).$$

For a given random sample  $X_1, X_2, \dots, X_n$

$$\hat{\theta}_{ML} = \hat{\theta}_{ML}(X_1, X_2, \dots, X_n).$$

Remark

$X_1, X_2, \dots, X_n$  independent.

$$L(x_1, x_2, \dots, x_n; \theta) = \prod_{i=1}^n L(x_i; \theta)$$

$$L(x_i; \theta) = \begin{cases} p_x(x; \theta) \\ f_x(x; \theta). \end{cases}$$

Log-likelihood

$$\ln(L(x_1, x_2, \dots, x_n; \theta)) = \sum_{i=1}^n \ln(L(x_i; \theta))$$

↳ easier to maximize.

Ex Exponential( $\theta$ )  $f(x; \theta) = \begin{cases} \theta e^{-\theta x} & x \geq 0 \\ 0 & \text{otherwise} \end{cases}$

Likelihood

$$L(x_1, x_2, \dots, x_n; \theta) = \prod_{i=1}^n \theta e^{-\theta x_i}, \quad x_i \geq 0.$$
$$= \theta^n e^{-\theta \sum_{i=1}^n x_i}$$

Take ln.

$$\ln(L(x_1, \dots, x_n; \theta)) = n \ln \theta - \theta \sum_{i=1}^n x_i$$

Maximize  $\Rightarrow$  take derivative.

$$\frac{\partial}{\partial \theta} \ln(L(x_1, \dots, x_n; \theta)) = \frac{n}{\theta} - \sum_{i=1}^n x_i = 0.$$

Solve for  $\theta$ .

$$\hat{\theta}_{ML} = \frac{1}{\frac{1}{n} \sum_{i=1}^n x_i} \quad \leftarrow \text{sample mean.}$$

# Works with multiple parameters

Ex  $X_1, X_2, \dots, X_n \sim \mathcal{N}(\theta_1, \theta_2)$

↑ ↑  
unknown,

Likelihood

$$L(x_1, x_2, \dots, x_n; \theta_1, \theta_2) = \prod_{i=1}^n f(x_i; \theta_1, \theta_2)$$

$$= \prod_{i=1}^n \left( \frac{1}{\sqrt{2\pi\theta_2}} e^{-\frac{(x_i - \theta_1)^2}{2\theta_2}} \right)$$

$$\ln(L(x_1, \dots, x_n; \theta_1, \theta_2)) = \sum_{i=1}^n \left( -\ln(\sqrt{2\pi\theta_2}) - \frac{(x_i - \theta_1)^2}{2\theta_2} \right)$$

$$= -\frac{n}{2} \ln(2\pi\theta_2) - \frac{1}{2\theta_2} \sum_{i=1}^n (x_i - \theta_1)^2$$

Take derivatives.

$$\frac{\partial}{\partial \theta_1} : \frac{1}{\theta_2} \sum_{i=1}^n (x_i - \theta_1) = 0$$

$$\frac{\partial}{\partial \theta_2} : -\frac{n}{2\theta_2} + \frac{1}{2\theta_2^2} \sum_{i=1}^n (x_i - \theta_1)^2 = 0.$$



System of eq  $m$   $\theta_1, \theta_2$

$$\Rightarrow \hat{\theta}_1 = \frac{1}{n} \sum_{i=1}^n x_i$$

$$\hat{\theta}_2 = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\theta}_1)^2 \quad \uparrow \text{Biased}$$

Properties of MLEs.

Let  $X_1, X_2, \dots, X_n$  be a random sample,  
for a distribution w/ parameter  $\theta$ .

$$\hat{\theta}_{ML} = \hat{\theta}_{ML}(X_1, X_2, \dots, X_n)$$

①  $\hat{\theta}_{ML}$  is consistent.

$$\hat{\theta}_{ML} \xrightarrow[n \rightarrow \infty]{i.p} \theta$$

② Asymptotically unbiased

$$\lim_{n \rightarrow \infty} B(\hat{\theta}_{ML}) = 0,$$

③ CLT

$$Z = \frac{\hat{\theta}_{ML} - \theta}{\sqrt{\text{Var}(\hat{\theta}_{ML})}} \xrightarrow{d} N(0, 1)$$